

# Journal of Clinical Sleep Medicine

# **SCIENTIFIC INVESTIGATIONS**

# Characterizing Sleep Structure Using the Hypnogram

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**Objectives**: Research on the effects of sleep-disordered breathing (SDB) on sleep structure has traditionally been based on composite sleep-stage summaries. The primary objective of this investigation was to demonstrate the utility of log-linear and multistate analysis of the sleep hypnogram in evaluating differences in nocturnal sleep structure in subjects with and without SDB.

**Methods**: A community-based sample of middle-aged and older adults with and without SDB matched on age, sex, race, and body mass index was identified from the Sleep Heart Health Study. Sleep was assessed with home polysomnography and categorized into rapid eye movement (REM) and non-REM (NREM) sleep. Log-linear and multistate survival analysis models were used to quantify the frequency and hazard rates of transitioning, respectively, between wakefulness, NREM, and REM sleep.

**Results**: Whereas composite sleep-stage summaries were similar between the two groups, subjects with SDB had higher frequencies and hazard rates for transitioning between the three states. Specifically, log-linear models showed that subjects with SDB had more wake-to-

NREM sleep and NREM sleep-to-wake transitions, compared with subjects without SDB. Multistate survival models revealed that subjects with SDB transitioned more quickly from wake-to-NREM sleep and NREM sleep-to-wake than did subjects without SDB.

**Conclusions**: The description of sleep continuity with log-linear and multistate analysis of the sleep hypnogram suggests that such methods can identify differences in sleep structure that are not evident with conventional sleep-stage summaries. Detailed characterization of nocturnal sleep evolution with event history methods provides additional means for testing hypotheses on how specific conditions impact sleep continuity and whether sleep disruption is associated with adverse health outcomes.

**Keywords**: Sleep disruption, sleep-disordered breathing, sleep structure and event history modeling

**Citation:** Swihart BJ; Caffo B; Bandeen-Roche K; Punjabi NM. Characterizing sleep structure using the hypnogram. *J Clin Sleep Med* 2008;4(4):XXX-XXX.

uantifying sleep fragmentation is central in assessment of sleep quality. Traditionally, measures such as the arousal frequency and sleep-stage percentages have been used to appraise sleep quality in research and clinical practice. Although conventional metrics of sleep structure have provided useful insight into the biology of sleep, these parameters explain only part of the variance in outcomes such as daytime sleepiness associated with conditions that fragment sleep such as sleep-disordered breathing (SDB). Furthermore, many of the conventional measures provide only an overall summary of the entire night and unable to capture the temporal evolution of overnight events, the frequency of sleep-stage transitions, and the time between these transitions. Given the remarkable progress in our understanding of the neurobiology of the sleep-wake switch<sup>4</sup>

#### **Disclosure Statement**

This was not an industry supported study. The authors have indicated no financial conflicts of interest.

# Submitted for publication July, 2006 Accepted for publication March, 2008

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Journal of Clinical Sleep Medicine, Vol. 4, No. 4, 2008

and the underlying neural circuitry responsible for transitioning between rapid eye movement (REM) and non-REM (NREM) sleep,<sup>5</sup> adequately characterizing sleep-stage transitions is a priority to better define the influence of specific factors (e.g., age and sex) on normal sleep structure and organization. In addition, a careful portrayal of sleep-stage transitions is essential in clarifying the putative mechanisms through which conditions such as SDB mediate adverse health outcomes.

Several techniques have been used to derive measures of sleep quality that complement the repertoire of traditional metrics. Power spectral analysis of the sleep electroencephalogram (EEG),6 sleep spectrograms based on cardiopulmonary coupling,<sup>7</sup> and visual identification of cyclical alternating patterns<sup>8</sup> in the sleep EEG have revealed clinically meaningful changes in the sleep structure in health and disease. Although these techniques provide unique insight into sleep continuity, their use requires specialized expertise along with an appreciation of the associated limitations. With improvements in digital technology, many of aforementioned techniques are automated and being increasingly incorporated in commercially available sleep-scoring software.9 A relatively underutilized, but universally available, method for assessing sleep continuity is the hypnogram. The graphic representation of sleep-stage sequence across the night provides a visual depiction of the normal ultradian cycling of sleep. While the hypnogram provides a







qualitatively description of sleep structure, quantitative metrics based on the hypnogram are not as commonly used in research or clinical practice as are other measures such as the frequency of arousals. Visual scoring of arousals is labor intensive, time consuming, and fraught with low to modest interscorer and intrascorer reliability. Even when coupled with the distribution of sleep-stage amounts, the frequency of arousals is unable to characterize the full extent of information embedded within the hypnogram. It is certainly plausible that a clinical disorder increases the frequency of sleep-stage transitions but has no material impact on the total amount of time spent in each stage or perhaps even the number of arousals. Tabulating the number of sleep-stage shifts can be helpful<sup>10,11</sup> but is insufficient because it describes only one dimension of the hypnogram (i.e., number of shifts) while neglecting another (i.e., the time spent in a sleep stage before transitioning). Methods to describe temporal histories as depicted in the hypnogram are common in epidemiologic studies but have had limited application in sleep medicine. Although event-history models have been previously used in the context of examining determinants of sleep latency, such methods have not been employed in assessing the sleep-stage transitions and quantifying the impact of SDB on the nocturnal sleep structure. 12-16 Thus, the primary objective of the current investigation was to determine whether event-history models are able to quantify sleep fragmentation using the overnight hypnogram. Specifically, log-linear models and multistate survival analysis methods were used to model the number and rate of sleep-stage transitions, respectively, in a community sample of middle-aged and older adults with and without SDB.

#### **METHODS**

## **Study Sample and Covariate Data**

The current investigation used data from the Sleep Heart Health Study (SHHS), a multicenter study on SDB, hypertension, and cardiovascular disease. To Subjects for the SHHS were recruited from ongoing cohort studies on cardiovascular and respiratory disease. Details regarding the design and methodology for recruiting and characterizing study subjects have been previously described. Approval for the study protocol was acquired from the institutional review board of each participating institution and informed consent was obtained from all subjects. The baseline visit included interviewer-administered questionnaires to assess prevalent medical comorbidities (e.g., hypertension and cardiovascular disease), smoking history, caffeine and alcohol consumption, race, sex, and age. Systolic and diastolic blood pressure, height, weight, and neck circumference were also obtained on the night of the polysomnogram.

To assess the independent effects of SDB on sleep structure, a matched subset of the SHHS cohort with and without SDB was selected for the current study. Subjects with moderate to severe SDB were identified as those with a respiratory disturbance index that exceeded the 90<sup>th</sup> percentile of the entire cohort (RDI  $\geq$  22.3 events/h). Subjects without SDB were identified as those with an RDI below the 25<sup>th</sup> percentile of the entire cohort (RDI < 1.33 events/h). Extremes of SDB severity were used to increase the likelihood of finding differences in conventional measures of sleep structure Confounding due to demographic

factors was minimized by matching subjects with and without SDB on age, sex, race, and body mass index (BMI). The limits imposed on age and BMI were such that no matching pair differed by more than 10 years (1 standard deviation of SHHS cohort) of age and 5 kg/m² (1 standard deviation) in BMI. Other exclusion criteria included prevalent cardiovascular disease, hypertension, chronic obstructive pulmonary disease, asthma, coronary heart disease, history of stroke, and current smoking. Despite having a baseline cohort of 6441 subjects, only 60 subject pairs with and without SDB (n =120) met the strict inclusion criteria outlined above and could be individually matched to each other.

# Polysomnography

An overnight sleep study in the subject's home was conducted for each subject using the Compumedics P-series recording system (Compumedics, Australia). The recording montage included the following physiologic recordings: EEG (C<sub>3</sub>-A<sub>2</sub> and C<sub>4</sub>-A<sub>1</sub>), right and left electroocculograms, single-lead electrocardiogram, chin electromyogram, measurement of abdominal and thoracic effort by impedance plethysmography, oxyhemoglobin saturation by pulse oximetry, airflow (oral-nasal thermistor), body position (by mercury gauge), and ambient light. All sleep recordings were sent to a centralized reading center for visual analysis. Details of polysomnographic equipment, hookup procedures, failure rates, scoring, and quality assurance have been previously described.<sup>18</sup>

Sleep-stage scoring was performed by trained technicians according to the published guidelines. 19 Apneas were identified if airflow was absent or nearly absent for at least 10 seconds. Hypopneas were identified if discernible reductions in airflow or thoracoabdominal movement (at least 30% below baseline values) occurred for at least 10 seconds. The RDI was defined as the number of apneas or hypopneas, each associated with a 4% decrease in oxygen saturation, per hour of sleep. Arousals were identified as abrupt shifts in the EEG frequency for at least 3-seconds. In REM sleep, scoring of arousals also required a concurrent increase inactivity of the chin electromyogram.<sup>20</sup> The arousal index was defined as the average number of arousals per hour of sleep. Conventional parameters of sleep structure included sleep latency, total sleep time, sleep efficiency (total sleep time/time in bed), and percentages of NREM and REM sleep. Subjects with visually identified poor-quality EEG were not eligible for the current analysis. Other exclusionary criteria included poor-quality oximetry or respiratory signals and inability to visually score sleep.

#### **Statistical Analysis**

To characterize nocturnal event histories in the sleep hypnogram, two distinct methods were employed: multistate survival analysis and log-linear models. Multistate survival models describe a finite number of states together with all possible transitions that can occur between those states. <sup>21,22</sup> The movement of subjects between states is governed by a set of transition rates (or hazard rates) that can be modeled using proportional hazards regression. <sup>23</sup> In the context of modeling overnight stage transitions, sleep was represented using three states: wake,









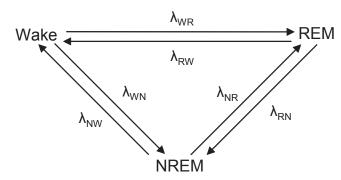


Figure 1—A schematic of the six possible transitions between wake, REM [rapid eye movement], and non-REM [NREM]). λpq is the hazard rate of making the transition from stage p to stage q.

NREM sleep, and REM sleep (Figure 1). The multistate variant<sup>24</sup> of the proportional hazards regression model that depicts the dynamics of sleep-stage transitions for subject i can be written as follows:

$$\log \lambda_s(t|x_i) = \log \lambda_{o(s)}(t) + \beta_s * x_i + \text{covariates}$$

Here, the strata index s indicates the type of sleep-stage transition (e.g. NREM-to-REM, Figure 1),  $\lambda_{o(s)}(t)$  is the distinct baseline hazard rate for each type of sleep-stage transition, x is an indicator variable for disease status (SDB versus no-SDB), and  $\beta_s$  is the regression coefficient for strata specific log(transition rate) comparing those with SDB compared to those without SDB..25 Due to the fact that a subject can cycle through all three sleep states several times during the night, six different types of transitions are distinguished, and each of these transitions can occur more than once. To estimate rates of transitioning in the multistate model, the data must be structured in a person-period format taking into account all possible competing transitions.<sup>26</sup> For example, a NREM sleep duration that transitions into REM sleep would be expanded to two data records: NREM-to-REM transition (observed record) and NREM-to-wake transition (censored record). The designation of the former as "observed" and the latter as "censored" indicates the occurrence of the NREM-to-REM transition during a period of risk for either transition. A stratified extension of proportional hazard models was fitted with the PHREG procedure in SAS (SAS Institute, Inc., Cary, NC). The robust sandwich variance estimator was used to account for intrasubject correlation, and ties were handled as proposed by Efron.<sup>27</sup> The stratified proportional hazards model was used because it can incorporate several states (e.g., wake, NREM, and REM) between which transitions may take place at distinct hazard rates. The STRATA specification of the PHREG procedure allows model fitting when the hazard functions across groups can be assumed to be parallel for a particular transition type but not across the different types of transitions. Thus, the stratified proportional hazards model accommodates the requirement that the baseline hazard rates for the six different transitions shown in Figure 1 are not necessarily similar.

To model the frequency of transitions as a function of group status, Poisson log-linear models were employed.<sup>28</sup> Poisson loglinear models, a specialized case of generalized linear models, are commonly used to model contingency tables. In the context of modeling the frequency of sleep-stage transitions, there

Table 1—Characteristics of Subjects With and Without Sleep-Disordered Breathing (SDB)

Variable	SDB	No SDB	p value <sup>a</sup>
Demographic			
Age, y	$62.7 \pm 10.8$	$62.3 \pm 10.6$	0.31
Male, no. (%)	38 (63.3)	38 (63.3)	1.00
White, no. (%)	52 (86.7)	52 (86.7)	1.00
BMI, kg/m <sup>2</sup>	$30.7 \pm 5.2$	$29.2 \pm 4.5$	< 0.0001
Polysomnographic			
RDI, events/h	$34.0 \pm 12.1$	$0.63 \pm 0.4$	< 0.0001
Total sleep time, min	$353.3 \pm 59.5$	$362.9 \pm 56.3$	0.38
Sleep latency, minb	$20.5 \pm 15.7$	$22.1 \pm 18.5$	0.69
Sleep efficiency, %	$81.9 \pm 10.3$	$83.0 \pm 9.2$	0.98
Sleep stage, %c			
1	$6.5 \pm 4.5$	$5.7 \pm 3.5$	0.21
2	$58.6 \pm 10.5$	$57.3 \pm 11.8$	0.52
SWS	$15.6 \pm 11.6$	$16.0 \pm 12.5$	0.86
REM	$19.2 \pm 7.2$	$21.0 \pm 5.6$	0.09
Arousal frequency,			
events/h	$28.0 \pm 12.0$	$13.7 \pm 5.7$	< 0.0001

Data are presented as mean  $\pm$  SD unless otherwise indicated. <sup>a</sup>Group differences by sleep-disordered breathing (SDB) status were determined by the Wilcoxon signed-rank test for categorical

<sup>b</sup>Denotes the latency to the first onset of sleep.

variables and paired t test for continuous variables.

<sup>c</sup>Stage 1, stage 2, slow-wave (SWS), and rapid eye movement (REM) sleep are expressed in percentage of total sleep time. SWS represents the combination of sleep stages 3 and 4.

are two distinct groups that can each repeatedly experience six possible transition types. The basic concept of the log-linear modeling involves fitting a model to the observed frequencies contained within the  $2 \times 6$  contingency table. The model is parameterized for row and column effects as follows:

$$\log (F_{\alpha}) = \mu + \varphi^{G} + \varphi^{S} + \varphi^{GS} + \text{covariates}$$

 $\begin{array}{l} log \; (F_{\it ab}) = \mu + \phi_{\it a}^{\; \rm G} + \phi_{\it b}^{\; \rm S} + \phi_{\it ab}^{\; \rm GS} + covariates \\ In the above equation, log(F_{\it ab}) is the log of the expected cell \end{array}$ frequency for cell ab in the contingency table; µ is an intercept (the referent cell's mean natural log of expected frequency);  $\varphi_a^G$ and  $\varphi_b^S$  represent the main effects of group status and transitiontype, respectively; and  $\phi_{ab}^{GS}$  estimates the interaction between group and transition-type effects. Generalized estimating equations were used to account for the interdependence among the cells.<sup>29,30</sup> As opposed to the multistate approach, which models events over time and accounts for censoring, the structure of the data for log-linear analysis is only concerned with the number of the transitions observed. The coefficients produced by this model and linear combinations thereof were appropriately transformed to render estimates of relative frequencies of particular sleep-stage transition types as a function of group status (SDB versus no SDB). The log-linear analysis was conducted using the GENMOD procedure in SAS. Both the multistate and log-linear models included the matching variables age, sex, race, and BMI. All p values are for 2-sided tests.

#### **RESULTS**

A matched sample of 60 subjects with and without SDB was identified from the SHHS cohort. As expected, the two groups were similar with respect to age, sex, and race (Table 1). However, a small but statistically significant difference was noted





Table 2—Results from Log-Linear and Multistate Models for Sleep-Stage Transitions in SDB for the Entire Night

Sleep-stage	Log-linear analysis					Multistate analysis		
transition	Frequency of transitions <sup>s</sup>		RRb			$HR^b$		
	SDB	No SDB	RR	95% CI	p value	HR	95% CI	p value
Wake $\rightarrow$ NREM	1725	1368	1.26	1.07, 1.48	0.005	1.10	1.02, 1.20	0.02
$NREM \rightarrow Wake$	1579	1200	1.32	1.11, 1.56	0.001	1.50	1.30, 1.74	< 0.0001
$NREM \rightarrow REM$	346	351	1.02	0.81, 1.29	0.85	1.04	0.68, 1.57	0.87
$REM \rightarrow Wake$	358	324	1.17	0.91, 1.50	0.21	2.26	1.51, 3.40	< 0.0001
$REM \rightarrow NREM$	160	134	1.20	0.90, 1.59	0.20	1.57	0.89, 2.77	0.12
Wake $\rightarrow$ REM	175	114	1.42	1.02, 1.96	0.04	0.86	0.49, 1.50	0.60

<sup>&</sup>lt;sup>a</sup> Unadjusted ratios comparing subjects with sleep-disordered breathing (SDB) to those without SDB (No SDB).

in the BMI between subjects with and without SDB. Restricting the matching limits in BMI to less than 5 kg/m² or age to less than 10 years to improve the degree of matching led to a significant decrease in the overall sample size. Thus, BMI, age, and other matching covariates were included in all multivariable models. Subjects with SDB had a mean RDI of 34.0 events per hour (median: 30.6, interquartile range: 26.4-39.1), whereas those without SDB had a mean RDI of 0.63 events per hour (median: 0.67, interquartile range: 0.32-0.91). As expected, the overall arousal frequency was higher in SDB subjects, compared with healthy subjects (no SDB). Surprisingly, despite obviously large differences in disease severity (i.e., RDI), conventional measures of sleep structure such as total sleep time, percentage of total sleep time in each sleep stage, and sleep efficiency were similar between the two groups (Table 1).

The transition frequencies of wake-to-NREM sleep and NREM sleep-to-wake were significantly higher in subjects with SDB (Table 2). Log-linear models revealed that SDB conferred a 26% and 32% increase in propensity of wake-to-NREM sleep and NREM sleep-to-wake transitions, respectively. The higher relative frequency of these two transition types suggests that SDB can disrupt sleep continuity with oscillations between NREM sleep and wakefulness. The log-linear model also showed that SDB increases the number of transitions from wake-to-REM sleep. Multistate models examining the hazards of each sleep-stage transition type revealed that there was an increase in the rate of wake-to-NREM sleep and NREM sleepto-wake transitions (Table 2), confirming the findings of the log-linear analysis. The adjusted hazard rate ratios of wake-to-NREM sleep and NREM sleep-to-wake transitions were 1.10 (95% confidence interval: 1.02, 1.20) and 1.50 (95% confidence interval: 1.30, 1.74), respectively. In addition, multistate models demonstrated an increase in the rate for REM sleep-towake, depicting those with SDB having a 2.26 (95% confidence interval: 1.51, 3.40) times greater likelihood of transitioning from REM sleep to wakefulness than those without SDB.

Recognizing the overnight heterogeneity in the distribution of NREM and REM sleep and in the frequency of sleep-stage transitions over the course of the night, analyses were performed dividing each subject's total sleep period into two segments. Accounting for differences in total sleep time, the first and second segments of sleep were determined as the first half and second half, respectively, of each individual's total sleep time, as opposed to using a global arbitrary cutpoint. Log-linear

and multistate models were then reconstructed for each segment of the sleep period (Table 3). Analyses by segment of the night showed that those with SDB had higher rates of awakening from both NREM and REM sleep in both segments. Transition frequencies for wake-to-NREM sleep and NREM sleep-to-wake were significantly higher in SDB subjects regardless of segment of night, confirming the overall-night results. It was also found that subjects with SDB had a higher transition frequency for wake-to-REM sleep and REM sleep-to-wake in the first but not in the second half of the sleep period. Finally, multistate analyses showed that, compared with the first half of the sleep period, subjects with SDB also displayed a greater propensity for sleep fragmentation, with significantly higher rates of wake-to-NREM sleep and REM-to-NREM transitions in the second half of the sleep period.

# DISCUSSION

The primary objective of the current study was to investigate whether characterizing sleep through log-linear and multistate analyses would demonstrate differences in sleep structure between those subjects with and without SDB. Using a matched sample of middle-aged and older adults recruited from the general community, the current investigation showed that, in the absence of confounding medical conditions, conventional measures of sleep structure were similar between those with and without SDB despite being at the extremes of health and disease, respectively. In contrast, log-linear and multistate models showed notable differences in sleep structure between the two groups. Subjects with SDB were noted to have a greater number and higher rates of sleep-stage transitions than those without SDB, suggesting that the occurrence of apneas and hypopneas during sleep can alter the duration spent in distinct sleep stages throughout the night without altering the overall summaries of sleep-stage amounts or total sleep time. In addition, the present study also showed that the propensity to transition from one stage to another is different between the first and second half of the sleep period.

The finding that SDB disrupts sleep continuity is not unexpected. It is well established that partial or complete collapse of the upper airway during sleep and the ensuing apneas and hypopneas often terminate with a brief EEG arousal. Recurrent arousals lead to state instability with recurrent to-and-fro transitions between different sleep stages. Event-related arousals from sleep are not all equal because there is much interin-







<sup>&</sup>lt;sup>b</sup>Results are adjusted for age, sex, race, and body mass index. RR refers to relative ratio and HR refers to hazard ratio; CI, confidence interval; REM, rapid eye movement NREM, non-REM.



Table 3—Results from log-linear and multistate models for sleep-stage transitions by segments of night

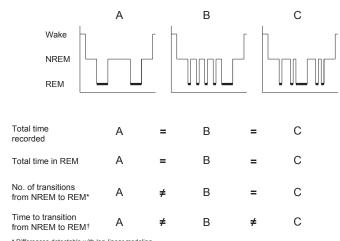
Sleep-stage transition		Log-linear analysis				Multistate analysis HR <sup>b</sup>		
Frequency of Transitions <sup>a</sup>			RRb					
	SDB	No SDB	RR	95% CI	p value	HR	95% CI	p value
First segment of nigh	t							
Wake $\rightarrow$ NREM	764	611	1.26	1.07, 1.48	0.005	1.09	0.99, 1.20	0.08
$NREM \rightarrow Wake$	712	553	1.30	1.10, 1.54	0.002	1.46	1.20, 1.78	0.0002
$NREM \rightarrow REM$	145	149	1.07	0.87, 1.32	0.53	1.16	0.71, 1.90	0.56
$REM \rightarrow Wake$	125	96	1.38	1.05, 1.81	0.02	2.77	1.55, 4.93	0.001
$REM \rightarrow NREM$	69	72	1.05	0.80, 1.37	0.73	1.05	0.56, 1.96	0.88
Wake $\rightarrow$ REM	63	32	1.87	1.29, 2.72	0.001	1.05	0.37, 3.00	0.92
Second segment of ni	ight							
Wake $\rightarrow$ NREM	961	757	1.26	1.02, 1.55	0.03	1.12	1.01, 1.24	0.03
$NREM \rightarrow Wake$	867	647	1.33	1.06, 1.67	0.01	1.53	1.27, 1.86	< 0.0001
$NREM \rightarrow REM$	201	202	1.01	0.76, 1.33	0.96	0.96	0.58, 1.62	0.89
$REM \rightarrow Wake$	233	228	1.06	0.81, 1.37	0.68	2.08	1.29, 3.35	0.003
$REM \rightarrow NREM$	91	62	1.19	0.85, 1.68	0.31	2.37	1.12, 5.02	0.02
Wake $\rightarrow$ REM	112	82	1.15	0.82, 1.6	0.43	0.79	0.41, 1.52	0.48

<sup>&</sup>lt;sup>a</sup> Unadjusted ratios comparing subjects with sleep-disordered breathing (SDB) to those without SDB (No SDB).

dividual and intraindividual variation in whether an apnea or hypopnea leads to a shift between sleep stages or a transition to wakefulness. The results of the current study illustrate that conventional measures of sleep structure tend to collapse a temporally evolving process and limit the ability to reach inferences regarding secular trends across groups. Nevertheless, composite summary measures provide useful and necessary information because knowing if the total time spent in a stage is similar across groups assists in distinguishing whether sleep is actually more fragmented or the greater frequency of transitions is simply a result of differences in total sleep time. In addition to characterizing transition frequencies and rates, event history methods, as employed herein, also afford modeling of directionality of transitions. For example, a state change from wake-to-NREM sleep is clearly distinct from the transition of NREM sleep-to-wake. Moreover, the methods of multistate survival analysis allow the dynamic notion of competing risks to be applied to the evolution of sleep, in which a transition from one to any of the other states is possible. Such distinction of state evolution of sleep is not possible with sleep-stage percentages or arousal counts because markedly differing sleep profiles can be congruent on these measures. Even counting the number of sleep-stage transitions is insufficient because it does not describe the time or rate of a particular type of sleep-stage transition. As shown in Figure 2, the percentages of NREM and REM sleep across different hypnograms can be similar, whereas overt differences can exist in the number and the rate of sleep-stage transitions. Log-linear and multistate models quantify these differences that are often visually apparent in the hypnogram.

There are several strengths of this study that merit discussion. The exclusion of medical comorbidities and matching on race, sex, BMI, and age minimized the concern for confounding and permitted a thorough assessment of the independent effects of SDB on sleep structure. Given that SDB is commonly associated with substantial medical comorbidity, identifying a sample of matched subjects free of such conditions is a major strength. Furthermore, the finding of similar conventional

sleep-stage summaries in subjects with and without SDB in our sample highlights the importance of characterizing the temporal evolution of sleep with methods that capitalize on distinct dimensions (i.e., frequency and rate) of an event. Availability of sleep recordings in the home is an additional strength because home-based studies limit the potential of the "first-night" effect on sleep structure that is common with in-laboratory polysomnograms.<sup>31-33</sup> Finally, the use of a well-characterized



† Differences detectable with log-linear modeling † Differences detectable with multi-state survival analysis

Figure 2—Differences in results obtained from survival and log-linear analysis of sleep-stage transitions illustrated using 3 hypothetical hypnograms. Total time recorded and time in rapid eye movement (REM) sleep are equivalent among the three hypnograms, demonstrating the limits of information gained by relying on composite summary measures alone to capture differences in sleep structure. Number of transitions from non-REM (NREM) to REM allow for a quantitative distinction to be made between profiles A and B and between A and C but not between B and C. Time to transition from NREM to REM provides a fuller description of profiles, enabling a quantifiable distinction among the three hypnograms.







<sup>&</sup>lt;sup>b</sup>Results are adjusted for age, sex, race, and body mass index. RR refers to relative ratio and HR refers to hazard ratio; CI, confidence interval; REM, rapid eye movement NREM, non-REM.



cohort recruited from the general-community cohort minimizes potential biases that are often inherent in clinic-based samples.

Despite these strengths, the current study has several limitations. First, a simplified approach for the assessment of transition frequency and rate was used that did not include individual stages of NREM sleep. The decision to limit classification to wake, NREM sleep, and REM sleep was driven by the need to answer the question of whether event-history methods provided any additional insight into the macrostructure of sleep. In light of the robust findings from modeling the three states, further application of these methods to the assessment of sleep structure with individual NREM sleep stages represents a logical extension. Second, because both of the methods employed in our analyses are based on visually scored sleep stages, poor reliability of scoring could impact the derived inferences. As observed by others, scoring of stage 1 sleep was least reliable in the SHHS. However, if stages are recoded as wake, NREM sleep, and REM sleep, as was needed for the log-linear or multistate analyses, interscorer comparisons in the SHHS yield kappa statistics in the range of 0.87 to 0.90.34 Third, stratification by sleep-stage transition type and further by segments of the sleep period diminishes the power to detect differences, especially if particular transition types occur infrequently (e.g., wake-to-REM sleep). However, the current analyses set the stage for future work with the entire SHHS cohort that will provide sufficient power to identify potential determinants of sleep-stage transitions. Fourth, a distinct feature of the proportional hazards model is that it leaves the underlying hazard for a specific type of transition unspecified. Although this is a major strength of the proportional hazards model, knowing the hazard rate in the reference group is sometimes desirable. Implementation of parametric models can provide these reference hazards and represent yet another extension of the current work.<sup>35</sup> Finally, the methods proposed herein characterize the continuity of sleep using 30-second epochs and thus cannot fully depict events (e.g., arousals or period of microsleep) that occur within the epoch. Nonetheless, event-history methods can be applied to sleep stages that are scored using a shorter epoch period (e.g., 4 seconds) to better describe sleep microstructure.

The major implication of this study is that the characterization of sleep structure in SDB and other sleep disorders is better served by encompassing new quantitative characterizations along with the classic measures, particularly to aptly test hypotheses regarding the function and health-related effects of normal and abnormal sleep. Within this broader scope, understanding various dimensions of sleep continuity may carry significance with regard to predicting the relationship between sleep and general health. In light of findings of the current study, the newly suggested approaches of examining sleep structure could provide a more thorough understanding of how comorbidities affect sleep, as well as how normal sleep function, in turn, fulfills a crucial role in health and disease.

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#### **APPENDIX**

The following table illustrates the data expansion necessary for multistate survival analysis. On the left, a data record of 1 state change is shown. Individual *i* transitions from non-rapid eye movement (NREM)-to-rapid eye movement (REM) sleep. To convert this record to a person-period format, the record would be expanded to 2 records that reflect all possible transitions from NREM sleep. These include a NREM-to-REM transition (observed) and NREM-to-wake transition (censored).

id	Transition type	<b>Duration of</b>		id	Transition type	<b>Duration of</b>	Observed (1) or
		state, min				state, min	Censored (0)
i	$NREM \rightarrow REM$	24.3	$\rightarrow$	i	$NREM \rightarrow REM$	24.3	1
				i	$NREM \rightarrow Wake$	24.3	0





